Estimation of the Required Amount of Hydrological Exploration in Lignite Mining Areas on the Basis of Hypothetical Hydrogeological Models

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ABSTRACT

Mine drainage is a necessary but very costly precaution for open-pit lignite mining
in sandy aquifers. Consequently, the minimization of the number of drainage wells
and their optimal operation become important tasks in designing mine drainage sys-
tems. Comprehensive groundwater flow models have to be used, both, for the
design of drainage wells, and for the analysis of water management strategies in
mining areas . The accuracy of computations with such models depends on the pre-
cision of the underlying hydrogeological informations. In order to get these infor-
mations detailed and costly hydrogeological explorations have to be done in the
mining regions.

The basic informations are obtained using exploration drilling. The cost for hydro-
geological exploration are approximately a linear function of the number of
exploration bore holes. Therefore the reduction of drilling gets a key role in
reducing costs of exploration. This might be done by:

- increased use of geophysical exploration methods,
- complex analysis of exploration results using mathematical statistical
  methods,
- precise estimation of the required amount of hydrogeological informations.

The paper describes a mathematical approach to support the complex decision
making procedure of estimating the optimal amount of hydrogeological exploration
with respect to a given mine drainage goal.

INTRODUCTION

For 1985 in the German Democratic Republic (GDR) an annual lignite production of 300 *Mill.tons / annum* is planned. The principal mining technology is open-pit mining. The lignite seams are embedded in quarternary/ tertiary aquifer systems. These aquifer systems have to be drained to satisfy the geomechanical stability of the slopes of the open-pit mines. In 1984 about 1.7 *Bill.* m^3 mine drainage water has been pumped out, operating more then 7000 drainage wells. Therefore, approximately 17 % of the total mining cost are required [1]. Consequently, the minimization of the number of drainage wells and their optimal operation become important tasks in designing mine drainage systems.

The extensive mine drainage causes manifold impacts on the water resources in mining areas and significant conflicts between different water users, [2]. Groundwater flow models have to be used, both, for the design of drainage wells, and for the analysis of water management strategies in mining areas [3].

The accuracy of computations with comprehensive groundwater flow models depends on the precision of the underlying hydrogeological informations. In order to get these informations detailed and costly hydrogeological explorations have to be done. Generally, hydrogeological exploration is based on the following techniques:

- *exploration drilling* for exploration including the collection and analysis of samples of the material in the bore hole, resulting in point informations on the hydrogeological structure,

- *pumping tests* for the estimation of transmissivities and specific yields being representative for a small region of the aquifer (a few 100 m^2),

- *geophysical methods* to get detailed informations within bore holes, and above all using surface methods and remote sensing techniques in order to obtain local and regional informations on the geohydrological system.

The basic informations are obtained using *exploration drilling*. The cost for hydrogeological exploration are approximately a linear function of the number of exploration bore holes. Therefore, the reduction of drilling gets a key role in reducing cost of exploration. This might be done by:

- increased use of geophysical exploration methods,

- complex analysis of exploration results using mathematical statistical methods,

- precise estimation of the required amount of hydrogeological informations.

In the paper we will concentrate on the last mentioned alternative.

At present, in the GDR the following tools are used for the *estimation of the required amount of hydrogeological exploration* :

a) standard values in the field of lignite exploration according to several stages of exploration and according to the type of the coal seam and deposit [4], see Figure 1a.

b) catalogue of groundwater deposits [5]; the parameters of the deposits are characterized by statistical values (mean, dispersion, variance). Assuming a required precision of exploration, the required amount of exploration can be estimated, see Figure 1b.

Both the tools give only rough estimates and they do not consider the aim of exploration, its role in the complex economic system of *exploration - mine drainage - mining*. Principally, the amount of exploration is estimated by experts taking into the account this complexity, but more or less on a subjective basis, see Figure 1c.

Obviously, the objectives to minimize mine drainage cost (e.g. minimizing the number of drainage wells) and to minimize exploration cost (e.g. minimizing the number of exploration drillings) are contradictory. - The less hydrogeological
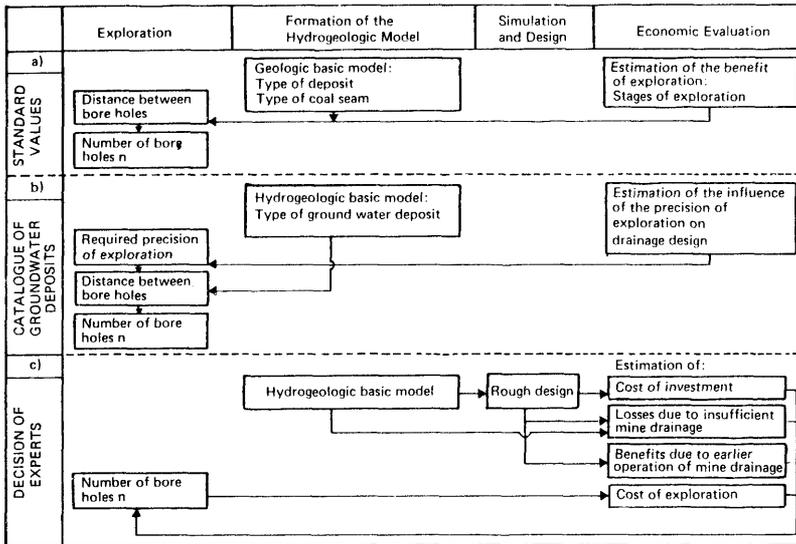
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Figure 1: Methods for the estimation of the required amount of hydrogeological exploration in the GDR

exploration is done, the more uncertain are the hydrogeological parameters for drainage well design. In the case of uncertain parameters drainage wells are more or less overdimensioned or the risk of geomechanical averages increases. Nevertheless, detailed hydrogeological exploration may be omitted if the risk of averages (and its economic consequences) is small in comparison with the benefits of saved exploration capacity and earlier operation of the drainage system [6].

In the following we propose a mathematical approach to support the complex decision making procedure of estimating the amount of hydrogeological exploration, see also [7].

METHODOLOGICAL APPROACH

Hydrogeological exploration in lignite mining areas aims above all at the estimation of hydrogeological parameters for calculations of the groundwater flow caused by mine drainage in the aquifer system. Such calculations include:

- estimation of flow to dewatering wells as the basis of mine drainage design,
- estimation of the total discharge of the mine drainage system for the design of mine water treatments plants, and for water management decisions,
- estimation of groundwater table variations (both, lowering in drained areas, and rise in abandoned mining areas) for water management and environmental decisions.

Generally, the installation and operation of mine drainage systems becomes the most costly task for water related decisions in mining areas. For the design of mine drainage systems the most detailed hydrogeological informations, their highest accuracy is needed. Consequently it is reasonable to assume, that the *amount of*

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hydrogeological exploration is estimated depending on the required exploration for the design of the mine drainage system. For an extention see Section 4.

Our approach is based on the following principal presumptions:

- For the area under consideration tentative geological and hydrogeological investigations have been done including preliminary (rough) exploration.
- The objectives of drainage are defined (location and depth of the lignite mine) and the drainage system is drafted (type of the drainage system, its depth and location).

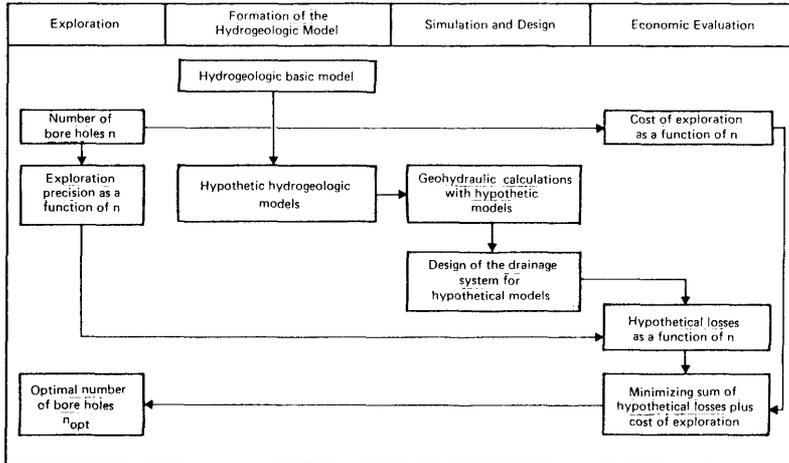The working steps of the proposed approach are summerized in Figure 2.

Figure 2: Working steps

Hydrogeological Schematization and Parameter Model

The complicated and partly random structure of hydrogeolocial systems necessitates their schematization because:

- hydrogeological exploration results either in rough estimates of the structure (geophysical methods) or in random samples (drilling, pumping tests),
- calculations of the groundwater flow system require a schematized and more or less simplified hydrogeolocial model of the real system.

Let us assume that the principle vertical structure (stratification) of the hydrogeological system under study is known. Usually this knowledge is well based on general geological informations and preliminary explorations. From that we obtain a vertical schematization. Define $\mathbf{P} = (p^{(1)}, p^{(2)}, \cdots)^T$ the set of hydrogeological parameters needed to quantify this schematization. The horizontal parameter distribution is called *parameter function* $f_P(x,y)$.

In the course of hydrogeological exploration (exploration drilling) point samples of the parameter function are obtained.

$$P(x_j, y_j) = f_P(x_j, y_j) \quad , \quad j = 1, \ldots, n_b \qquad (1)$$

with $\quad n_b$ – number of bore holes

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Based on these point samples an approximative parameter function has to be estimated.

$$\tilde{P} = \tilde{f}_P(x,y) \approx f_P(x,y) \tag{2}$$

This is a widely studied and discussed problem in hydrogeological research and practice. From our point of view only two concepts are practically important:

- the *kriging interpolation technique* of best linear unbiased estimation of a regionalized variable, for a review see [8],
- the concept of *geohydraulically representative parameters*.

The kriging interpolation technique and others are used to estimate *local* parameters, whereas the second concept results in *regional* representative parameters. Our approach is based on the second concept.

The area under consideration Σ (area being influenced by the mining) should be divided into N subareas $\Delta\Sigma_i$.

$$\Sigma = \Sigma_1 \cup \Sigma_2 \cup \cdots \cup \Sigma_N \tag{3}$$

Each subarea is characterized by a uniform stratification and an accompanying set of hydrogeological parameters. For the purpose of simplification we consider in the following only one parameter p_i for each subarea i, being the decisive parameter for drainage design. The approach can be extended for several parameters without principle difficulties.

$$p_i = p_i(x,y) \quad (x,y)\ \varepsilon\ \Delta\Sigma_i \quad,\quad i=1,\ldots,N \tag{4}$$

The parameter distribution within the subarea is assumed to be random, and the subarea should appertain statistically to a basic totality with regard to the p_i feature.

The results of hydrogeological explorations are used to estimate empirical distribution functions of the parameter p_i. E.g. statistical investigations of the empirical distribution functions of the hydraulic conductivity for a large number of aquifers and subareas have revealed different distributions ranging from the Normal distribution to a logarithmic bell-shaped distribution [9].

For geohydraulic calculations, the random parameter p_i has to be replaced by a constant *geohydraulically representative parameter* $p_{R,i}$. According to [9], the parameter $p_{R,i}$ is *geohydraulically representative*, if computations of the flow system (aimed at the design of the drainage system) with the representative value are adequate to computations with the real hydrogeological parameter $p_i(x,y)$ as a function in space. This means, the real hydrogeological system in the subarea is replaced by a schematized horizontal stratified subarea being homogeneously within each stratum.

Frequently, as an estimate of the geohydraulically representative value the arithmetic mean $\bar{p}_i$ from all samples in the subarea is used, assuming normal distributed parameters. For the parameter holds:

$$p_i : N(\bar{p}_i\ ,\ \sigma_i) \tag{5}$$

We need an estimate $\tilde{p}_i$ of the mean. Because $\tilde{p}_i$ is normal distributed, the estimate of the mean will be normal distributed:

$$\tilde{p}_i : N\left(\bar{p}_i\ ,\ \frac{\sigma_i}{\sqrt{n_{b,i}}}\right) \tag{6a}$$

with σ_i — variance

$n_{b,i}$ — number of samples (bore holes).

Under the given assumptions a confidence interval may be estimated as defined below.

$$\bar{p}_i - \frac{s_i}{\sqrt{n_{b,i}}} \cdot t_{a,m_i} < \tilde{p}_i < \bar{p}_i + \frac{s_i}{\sqrt{n_{b,i}}} \cdot t_{a,m_i} \tag{6b}$$

with α – error probability
  t_{a,m_i} – parameter of the t-distribution for unilateral questioning
  m_i – degrees of freedom $= n_{b,i} - 1$
  s_i – estimate of the variance.

A generalized method for the estimation of geohydraulically representative parameters (hydraulic conductivity, transmissivity) has been developed by [9]. The method although renders the estimation of the confidence of the calculated parameters. For the geohydraulically representative parameters holds:

$$p_{R,i} = \bar{p}_i \cdot c(v_i, \kappa) \tag{7a}$$

and for the confidence limits we obtain:

$$p_{R,i} - \frac{s_i}{\sqrt{n_{b,i}}} \cdot t_{a,m_i} \cdot c < \tilde{p}_i < p_{R,i} + \frac{s_i}{\sqrt{n_{b,i}}} \cdot t_{a,m_i} \cdot c \tag{7b}$$

with c – reduction factor
   v – variation coefficient
   κ – transformation coefficient

Both the simple arithmetic mean and more sophisticated parameter models for geohydraulically representative parameters open the possibility of confidence estimates.

$$p_{R,i} - \Delta p_{R,i}(n_{b,i}) \leq \tilde{p}_i \leq p_{R,i} + \Delta p_{R,i}(n_{b,i}) \tag{8}$$

With Eq. (8) we have a *relationship between:*

- the *exploration precision* in terms of the confidence of the hydrogeological parameter to be explored and
- the *amount of exploration* in terms of the number of exploration drillings.

Generally, the parameter model has to be chosen depending on the goal of exploration and on the expected hydrogeological situation.

Design of the Drainage System

The common technology for mine drainage in sandy aquifers are vertical dewatering wells (border and field well galleries). In general their design is based on uncertain data due to the random structure of the hydrogeological system and its restricted explorability. Furthermore, errors caused by schematizations and numerical calculations have to be taken into the account.

In open-pit mining the economic losses due to insufficient mine drainage (causing water inrushes) may be tremendous and may exceed the losses resulting from "over-design". The lower the risk of a breakdown of mine drainage the higher is the risk of uneconomic design. Nevertheless, the risk of damages due to "under-design" should be sufficiently small. Principally, the system of dewatering wells has to be designed in such a way that the technologically required groundwater depression within the mine will be satisfied with a certain reliability. Hence, the design depend on the reliability of the input data.

According to our principal presumptions we assume a fixed construction of drainage wells (drilling diameter, well screen and filter, capacity of pumps), as well as a given location of the dewatering galleries. Based on that, the number of wells, their distribution along the galleries become the decisive design values. These values depend strongly on the specific groundwater flow to the well gallery, the specific pumpage respectively, for a given groundwater depression target.

For the purpose of simplification the drainage well galleries should be divided in a finite number of sections Δg_j , $j=1,...,J$ with a constant specific pumpage q_j. Assuming a homogeneous horizontal hydrogeological structure along the gallery,

the well distance Δw_j for the section j will be constant and for the number of wells holds

$$n_{w,j} = \frac{\Delta s_j}{\Delta w_j} \qquad (9)$$

with $\qquad \Delta s_j$ – length of section Δg_j

For a given well construction and drainage target the well distance Δw_j is a multidimensional function, depending on the specific pumpage q_j and on the hydrogeological parameter p_j (the transmissivity) for the gallery section j. We assume steady state conditions.

$$\Delta w_j = f_w(p_j, q_j) \qquad (10)$$

Frequently the well distance for single aquifers is estimated from the following implicit mathematical model [10]:

$$q_j = T_j \cdot \frac{H_{g,j} - H_{w,j}}{\frac{\Delta w_j}{2 \cdot \pi} \cdot \ln \frac{\Delta w_j}{2 \cdot \pi \cdot r_{0,j}}} \qquad (11)$$

with $\quad q_j$ – specific pumpage for sector j,
$\quad\;\; T_j$ – actual transmissivity for sector j,
$\quad\;\; H_{g,j}$ – piecometric head at the drainage contour of sector j,
$\quad\;\; H_{w,j}$ – minimum piecometric head in the drainage wells,
$\quad\;\; r_{0,j}$ – hydraulic effective radius of the wells.

The specific pumpage is a function of the hydrogeological parameter of the subarea $\Delta\Sigma_i$.

$$q_j = f_q(p_1, p_2, \ldots, p_N) \qquad (12)$$

From Eq. (12), (10), (9) we would get

$$n_{w,j} = f_n(p_1, p_2, \ldots, p_N) \qquad (13)$$

Such a relationship would give us the interdependency between the reliability of the hydrogeological parameters to be explored and the design value of the drainage system. Unfortunately such a function can not be found for generalized problems. This is a consequence of the nonlinear, implicite well functions (e.g. Eq. (11)), and of the numerical solution of the groundwater flow problem. This will be discussed in the next section.

Geohydraulic Calculation with Hypothetic Hydrogeological Models

The specific pumpage q_j, $j=1,\ldots,J$ has to be estimated by the help of geohydraulic calculations. In general, for such calculations comprehensive system–descriptive groundwater flow models are used. Analytical solutions are rarely applicable. Consequently, their is no explicit relationship between the hydrogeological parameters of the system and the specific pumpage. Furthermore, we will not get an explicite relationship between random hydrogeological parameters and random specific pumpage (both in terms of a probability distribution function or as expectation values with confidence interval).

Empirical attempts to obtain empirical distributions (and confidence intervals) have been made using Monte-Carlo simulations, [11], or applying numerical approaches, [12]. The practical applicability is restricted due to the large amount of computations needed. In the following, we propose a reduced deterministic approach, based on the theory of statistical experiments [13].

Instead of a stochastic parameter model (Section 2.1) a discrete lumped parameter model is used. Each parameter is described as a finite number of possible realizations, chosen by the help of statistical methods.

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$$p_i : (p_i^{(1)}, p_i^{(2)}, \ldots, p_i^{(l)}) \tag{14}$$

Through simulation experiments with an appropriate groundwater flow model the effect of different realizations of the parameters - called *hypothetical hydrogeological model* - on the specific pumpage q_j and finally the design value Δw_j can be estimated.

If a linear relationship between the hydrogeological parameters p_i and the design value Δw_j holds true, two step experiments are needed. Two step experiments are also reasonable as a first approximation for nonlinear relationships, see below.

Each parameter is described by its expected lower and upper bounds:

$$\hat{p}_i : \left\{ \overline{\min p_i} \mid \overline{\max p_i} \right\} \tag{15}$$

These values may be estimated by previous hydrogeological explorations and practical experiences.

Define m the number of independent parameters to be explored. Consequently, 2^m experiments with hypothetical models are needed to cover all possible combinations of parameters. This might be with one parameter in m subareas or N parameters in m^{*} subareas with $m = N \cdot m^{*}$. Here the problem of numerical effort in case of more than ≈ 4 parameters becomes obvious (e.g. $m = 8$ results in 256 experiments).

To overcome this problem two alternatives should be investigated:

- application of simulation models with low computational effort,
- application of the theory of statistical experiments to reduce the number of experiments.

The outcome of the first alternative is restricted. In principle, the system-descriptive groundwater flow models with distributed parameters have to be used. That means, numerical models based on Finite-Differences-Methods or Finite-Elements-Methods are needed. A certain number of nodes (elements) is required to satisfy a acceptable accuracy. Furthermore, this number might increase with the number of parameters to be investigated.

In case of nonlinear relationships the deviation from the linear behavior can be estimated with an additional experiment for the central point $(\overline{\min p_i} + \overline{\max p_i})/2$. For strong nonlinearities more detailed computations are necessary, e.g. 3^m experiments in the case of a quadratic dependency.

From the experiments with the hypothetical models we obtain an evaluation to the effect of the hydrogeological parameter $\bar{p}_i$ in the subarea i on the drainage design, $n_{w,j}$, $j = 1, \ldots, J$. Obviously, the effect depends above all on the distance of the subarea from the location of the drainage system.

Define a hypothetical model $\mathbf{M}(\bar{p}_i)$ as a model characterized by a special parameter $\bar{p}_i$ for subarea i and any values according to Eq. (15) for the parameters of the other subareas l, $l = 1, \ldots, N$; $l \neq i$.

$$\mathbf{M}(\bar{p}_i) = \left\{ \hat{p}_1, \hat{p}_2, \ldots, \bar{p}_i, \ldots, \hat{p}_N \right\} \tag{16}$$

Analogously we define $\mathbf{M}(\bar{p}_i, \bar{p}_k)$ as

$$\mathbf{M}(\bar{p}_i, \bar{p}_k) = \left\{ \hat{p}_1, \hat{p}_2, \ldots, \bar{p}_i, \ldots, \bar{p}_k, \ldots, \hat{p}_N \right\} \tag{17}$$

In using this notation we can define *main effects* and *interaction effects*.

The *main effect* $\Delta n_{w,j}^i$ of the parameter $\bar{p}_i$ is defined as the difference between the mean for all parameter combinations for $\overline{\max p_i}$ minus the mean for all parameter combinations for $\overline{\min p_i}$.

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$$\Delta n_{w,j}^{i} = \Delta n_{w,j}(\bar{p}_i) = \frac{1}{K1}\left[\sum_1^{K1} \mathbf{M}(\overline{\max p_i}) - \sum_1^{K1} \mathbf{M}(\overline{\min p_i})\right] \tag{18}$$

with $K1$ - number of models ($=2^{N-1}$).

This relationship implies that the parameters of the subareas l ,$l=1,...,N$, $l \neq i$ are considered as a mean. Consequently for a linear dependency between $\Delta n_{w,j}$ and $\bar{p}_i$ $\Delta n_{w,j}$ is the number of wells for section j being additionally necessary if the parameter $\bar{p}_i$ changes from $\overline{\min p_i}$ to $\overline{\max p_i}$.

The *interaction effect* $\Delta n_{w,j}^{i,k}$ is defined as

$$\Delta n_{w,j}^{i,k} = \frac{1}{2}\left[\frac{1}{K2}\left[\sum_1^{K2} \mathbf{M}(\overline{\max p_i},\overline{\max p_k}) - \sum_1^{K2} \mathbf{M}(\overline{\min p_i},\overline{\max p_k})\right]\right.$$
$$\left. -\frac{1}{K2}\left[\sum_1^{K2} \mathbf{M}(\overline{\max p_i},\overline{\min p_k}) - \sum_1^{K2} \mathbf{M}(\overline{\min p_i},\overline{\min p_k})\right]\right] \tag{19}$$

with $K2$ - number of models ($=2^{N-2}$).

In the case of 3^m experiments or more a variance analysis has to be applied. The number of experiments can be reduced neglecting some of the interaction effects [13].

With Eq. (18) and (19) we have got a measure for the effect of parameter changes on the design values. As larger these main effects for the parameter $\bar{p}_i$ of a subarea $\Delta\Sigma_i$ are, as more important is the exploration in that subarea. Large interaction effects $\Delta n_{w,j}^{i,k}$ indicates, that the parameter $\bar{p}_i$ has a significant influence on the design with a special constellation of parameter $\bar{p}_k$. According to these numbers a step-wise exploration might be implemented. We will illustrate this in Section 3 for an example.

Economic Evaluation Depending on Exploration Precision

For the time being we will not consider the economic evaluation of the risk due to insufficient design of mine drainage, as explained above. We will base on the economic evaluation of the drainage design and of the exploration. In Figure 3 the principle economical relationships are depicted.

As lower the exploration precision is, as higher are the cost for mine drainage to be expected in order to minimize the risk of "under-design". On the other hand, the cost of exploration increases with the exploration precision.

Define $\mathrm{ex}C_d$ the cost of mine drainage being necessary for the given system it means the system would have been optimal designed with exact parameters. The additional cost for drainage design due to insufficient exploration precision we define as *hypothetical losses L*.

To develop a relationship between the exploration precision (the number of exploration drillings) and the hypothetical losses is rather difficult. The cost for mine drainage as well as the hypothetical losses depend linear on the number of exploration drillings. But the number of drillings is a nonlinear function of the parameters, see Eq. (11). As a first approximation we assume that the number of well drillings and consequently the hypothetical losses depend linear on the parameters $\tilde{p}_i$, $i=1,...,N$. This assumption can be checked estimating the number of exploration drillings for the lower and upper parameter bounds of Eq. (15) and for the mean parameter.

Consider one parameter $\tilde{p}_i$ with expected lower and upper bounds $\min\tilde{p}_i$, $\max\tilde{p}_i$. Define p_i^d its "pessimistic" value used for the design of the drainage system. Hypothetical losses due to parameter $\tilde{p}_i$ will equal zero, if the parameter p_i^d is equal to the real value $\tilde{p}_i$, and will reach the maximum, if $\tilde{p}_i$ is equal to the "optimistic" value p_i^o.
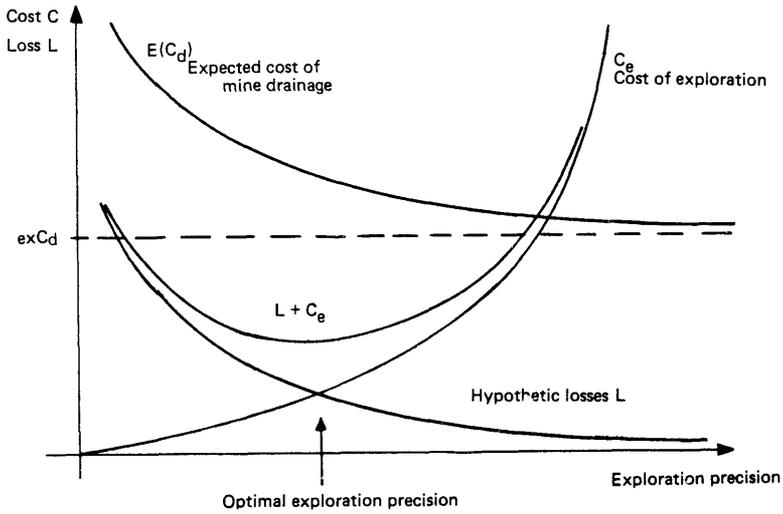
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Figure 3 : Economical evaluation

The losses are defined as

$$L(\tilde{p}_i) = \frac{\min C_d - \max C_d}{p_i^o - p_i^d} \cdot (\tilde{p}_i - p_i^d) \qquad (20)$$

with

$$p_i^d = \min\tilde{p}_i \ \rightarrow \ p_i^o = \max\tilde{p}_i$$
$$p_i^d = \max\tilde{p}_i \ \rightarrow \ p_i^o = \min\tilde{p}_i$$

or

$$L(\tilde{p}_i) = \frac{\max C_d - \min C_d}{\max\tilde{p}_i - \min\tilde{p}_i} \cdot |\tilde{p}_i - p_i^d| \qquad (21)$$

In Figure 4 the function is illustrated. For the parameter p_i^d we apply the generalized parameter model from Eq. (8).

$$p_i^d = \bar{p}_i \pm \Delta\bar{p}_i(n_{b,i}) \qquad (22)$$

That means, for the design either the lower or upper confidence limit (for a given probability) is used.

Inserting Eq. (22) into Eq. (21) we obtain

$$L(\tilde{p}_i) = \frac{\max C_d - \min C_d}{\max\tilde{p}_i - \min\tilde{p}_i} \cdot |\tilde{p}_i - \bar{p}_i \pm \Delta\bar{p}_i(n_{b,i})| \qquad (23)$$

For the normal distributed parameter model we get with Eq. (7)

$$L(\tilde{p}_i) = \frac{\max C_d - \min C_d}{\max\tilde{p}_i - \min\tilde{p}_i} \cdot |\tilde{p}_i - \bar{p}_i \pm \frac{s_i}{\sqrt{n_{b,i}}} \cdot t_{a,m_i}| \qquad (24)$$
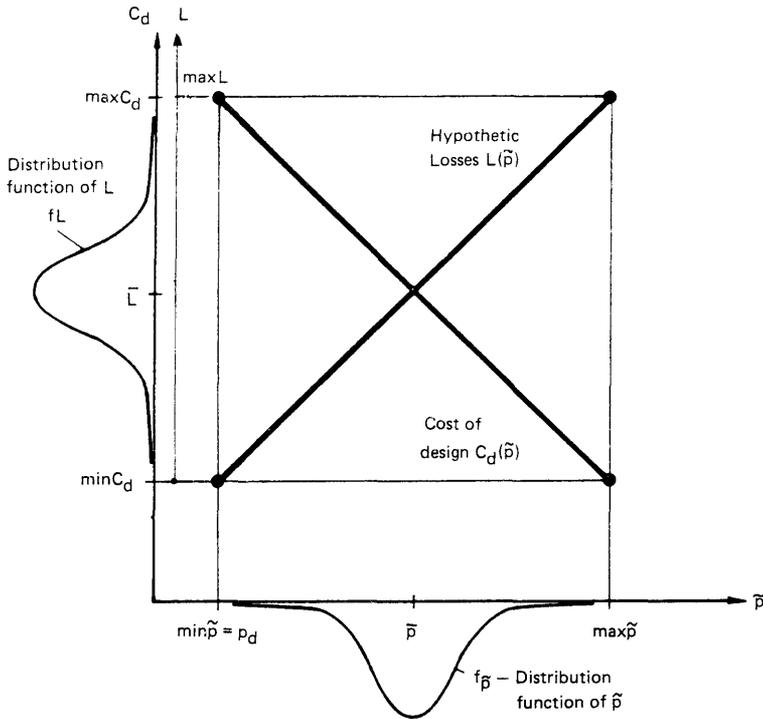
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Figure 4: Cost of design and hypothetical losses in relation to the parameter $\tilde{p}_i$ for $p_i^d = \min \tilde{p}_i$.

In Figure 5 this functional relationship is illustrated. Because $\tilde{p}_i$ is normal distributed, $L(\tilde{p}_i)$ is normal distributed with the mean

$$E\left[L(\tilde{p}_i)\right] = L(\bar{p}_i) \tag{25}$$

and the dispersion

$$\sigma_L^2 = (L(\bar{p}_i - \frac{s_i}{\sqrt{n_{b,i}}}) - L(\bar{p}_i))^2 \tag{26}$$

An open problem is the evaluation of the cost of drainage. Let c_w be the specific cost of one drainage well. Now we will replace the cost by using the *effects*, being defined in the previous section (Eq. (18), (19)).

$$\max C_d - \min C_d = |\ \Delta n_{w,j}^i\ | \cdot c_w \tag{27}$$

As explained above, $\Delta n_{w,j}^i$ is the number of wells being additional necessary, if the parameter $\tilde{p}_i$ changes from $\min \tilde{p}_i$ to $\max \tilde{p}_i$. - And this is the meaning of $\max C_d - \min C_d$!
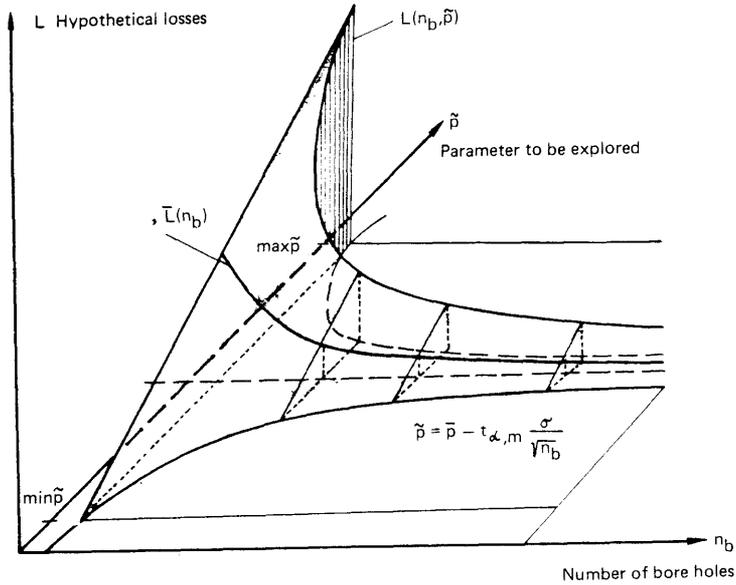
Inserting in Eq. (22) we obtain

Figure 5 : Hypothetical losses for normal distributed parameters

$$L(\widetilde{p}_i) = \frac{|\ \Delta n_{w,j}^i\ |\cdot c_w}{\max\widetilde{p}_i - \min\widetilde{p}_i}\cdot|\ \widetilde{p}_i - \bar{p}_i \pm \Delta\bar{p}_i\,(n_{b,i})| \qquad (28)$$

In the case of normal distributed parameters we can replace $\max\widetilde{p}_i - \min\widetilde{p}_i$

$$\frac{\max\widetilde{p}_i - \min\widetilde{p}_i}{2} = l\cdot s_i \qquad (29)$$

with s_i - variance

        l - measure of the quality of previous informations

Hence we get for Eq. (23) with Eq. (27)

$$L(\widetilde{p}_i) = \frac{|\ \Delta n_{w,j}^i\ |\cdot c_w}{2\cdot l\cdot s_i}\cdot|\ \widetilde{p}_i - \bar{p}_i \pm \frac{s_i}{\sqrt{n_{b,i}}}\cdot t_{a,m_i}| \qquad (30)$$

For the *cost of exploration* we assume a linear dependency on the number of bore holes. For instance, we might use the following simple expression

$$C_e = C_{e,c} + c_b\cdot\sum_{i=1}^{N} n_{b,i} \qquad (31)$$

with $C_{e,c}$ - constant cost of exploration

        c_b - specific cost of exploration (cost of one bore hole).

Optimal Exploration Precision

We obtain the optimal exploration precision in minimizing the sum of total hypothetical losses and cost of exploration. For a parallel exploration of all subareas holds:

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$$V = \sum_{i=1}^{N} L(\widetilde{p}_i) + C_{e,c} + c_b \cdot \sum_{i=1}^{N} n_{b,i} \rightarrow Min.! \tag{32}$$

The hypothetical losses $L(\widetilde{p}_i)$ depend on the number of exploration bore holes in the subarea i only (compare Eq. (28), (29)). Consequently, the minimum problem Eq. (32) can be separated in a set of minimum problems, without consideration of $C_{e,c}$. For each subarea a part of the constant cost $C_{e,c}$ may be added, as it is done in the example (Section 3.). But this does not change the optimal number of exploration drillings.

$$V_i = L(\widetilde{p}_i) + c_b \cdot n_{b,i} \rightarrow Min.! \quad , \quad i = 1, \ldots, N \tag{33}$$

With Eq. (28) we get

$$\frac{|\Delta n_{w,j}^i| \cdot c_w}{\max \widetilde{p}_i - \min \widetilde{p}_i} \cdot |\widetilde{p}_i - \bar{p}_i \pm \Delta \bar{p}_i(n_{b,i})| + c_b \cdot n_{b,i} \rightarrow Min.! \tag{34}$$

An exploration is only then reasonable, if it leads to more precise informations in comparison with the previous informations being fixed in $\max \widetilde{p}_i$, $\min \widetilde{p}_i$. For the minimum number of bore holes $\min n_{b,i}$ we obtain by using the confidence limit

$$\Delta \bar{p}(\min n_{b,i}) = \frac{\max \widetilde{p}_i - \min \widetilde{p}_i}{2} \tag{35}$$

From Eq. (35) $\min n_{b,i}$ may be obtained. Hence, Eq. (34) has to be investigated for

$$n_{b,i} > \min n_{b,i} \tag{36}$$

Eq. (34) does not describe a simple minimization problem due to the stochastic parameter $\widetilde{p}_i$. According to [14] the decision should be based on the minimization of expectation values (Bayes-decision). The expectation value for $\widetilde{p}_i$ is $\bar{p}_i$. Hence we obtain from Eq. (34) with Eq. (25) a deterministic problem being solvable.

$$\frac{|\Delta n_{w,j}^i| \cdot c_w}{\max \widetilde{p}_i - \min \widetilde{p}_i} \cdot \Delta p(n_{b,i}) + c_b \cdot n_{b,i} \rightarrow Min.! \tag{37}$$

If the function of losses is nonlinear, the expectation value of losses might not be obtained explicitly. Then the Monte-Carlo method should be used to estimate the expectation value of losses depending on the parameter model.

Finally lets consider the case of normal distributed parameters. For the minimum number of bore holes holds

$$t_{a,m_i} \cdot \frac{s_i}{\sqrt{\min n_{b,i}}} = \frac{\max \widetilde{p}_i - \min \widetilde{p}_i}{2} \tag{38}$$

With Eq. (29) we get

$$\frac{t_{a,m_i}}{\sqrt{\min n_{b,i}}} = \frac{l}{2} \tag{39}$$

$$\min n_{b,i} = \left(2 \cdot \frac{t_{a,m_i}}{l}\right)^2$$

t_{a,m_i} has to be taken from tables with

$$m_i = n_{b,i} - 1 \tag{40}$$

For $n_{b,i} > \min n_{b,i}$ we obtain from Eq. (37) and (30)

$$f(n_{b,i}) = \frac{|\Delta n_{w,j}^i| \cdot c_w}{2 \cdot l} \cdot \frac{t_{a,m_i}}{\sqrt{n_{b,i}}} + c_b \cdot n_{b,i} \rightarrow Min.! \tag{41}$$

Eq. (41) cannot be solved explicitely because t_{a,m_i} depends on $n_{b,i}$ (Eq. (40)). The simplest way for the solution is to calculate the values starting from $n_{b,i} = \min n_{b,i}$ until the minimum is reached.

To reduce the number of calculations, Eq. (41) might be solved for an estimate of $\overline{t_{a,m_i}}$.

$$\overline{n_{b,i}} \approx \left\{ \frac{1}{8} \frac{\Delta n_{w,j}^{i\,2} \cdot c_w}{l \cdot c_b} \cdot \overline{t_{a,m_i}} \right\}^{\frac{3}{2}} \tag{42}$$

The optimal $n_{b,i}$ has to be determined precisely in the vicinity of $\overline{n_{b,i}}$ with the help of Eq. (41).

TEST EXAMPLE

We will demonstrate the developed approach for a simplified example of the design of a mine drainage gallery. The schematized system is depicted in Figure 6.
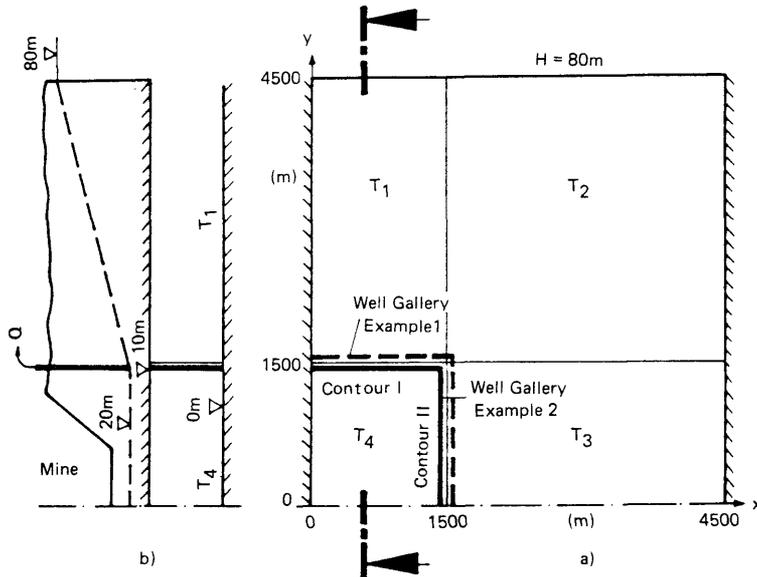
Figure 6 : Scheme of the test example, a) horizontal plan b) cross section

We consider a confined aquifer, schematized into 4 homogeneous subareas (with the transmissivities T_1, T_2, T_3, T_4). The variance of the real transmissivities within these subareas is expected to be small, the mean value $\overline{T}$ is approximately equal to the representative value T_R.

The well gallery has to be designed to satisfy a given groundwater depression next to the mine. Two alternatives of the location of the gallery are under consideration, as Figure 6 indicates (Example 1, Example 2). The piecometric head inside the drainage contour has to be equal 20 m.

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The transmissivities are expected to be in a given range.

$$2.3 \cdot 10^{-4} \leq T_1, T_2, T_3, T_4 \leq 2.3 \cdot 10^{-3} \quad m^2/sec. \tag{43}$$

The drainage gallery is divided into two sections (contour I, contour II). For these sections a uniform distribution of the wells is assumed. For more detailed investigations the drainage contour might be divided into more sections depending on the spatial variability of the groundwater flow (compare Figure 7, below).

For the geohydraulic calculations hypothetical hydrogeological models have been composed from the minimal and maximal expected transmissivities.

$$\min T_{1,\ldots,4} = 2.3 \cdot 10^{-4} \quad m^2/sec. \tag{44}$$

$$\max T_{1,\ldots,4} = 2.3 \cdot 10^{-3} \quad m^2/sec.$$

In Table 1 the hypothetical models are listed.

Table 1 : Hypothetical hydrogeological models for simulation experiments

Number	1	2	3	4	5	6	7	8	9	10	11	12	13	14	15	16
T_1	min	max	min	max	min	max	min	max	min	max	min	max	min	max	min	max
T_2	min	min	max	max	min	min	max	max	min	min	max	max	min	min	max	max
T_3	min	min	min	min	max	max	max	max	min	min	min	min	max	max	max	max
T_4	min	min	min	min	min	min	min	min	max	max	max	max	max	max	max	max

For each hypothetical model the specific pumpage q_I, q_{II} (steady state values) had to be estimated. We used a finite difference groundwater flow model [7]. The computational results are independent on the parameter T_4. That means, the results for the models 9−16 are the same as those for the models 1−8 (in the same order). Furthermore, the results for Example 1 and 2 are equal, assuming that the difference in the location of the gallery is negligible.

The computational result for the 8 experiments are illustrated in Figure 7. The specific pumpage was normalized by the maximum piecometric difference $\Delta H = 60\ m$.

For drainage contour I the maximum pumpage appertains to model 4 (maximum values for T'_1, T_2). As expected for contour II the maximum pumpage is reached for the models 7 and 8 (maximum values of T_2, T_3). The larger the pumpage is, the larger becomes its dependency on space.

Based on Eq. (11) and on mean values of the specific pumpage for the drainage contours I and II from the simulation experiments (Figure 7) the optimal well distance has been estimated using the following data:

$$H_c = 20\ m \qquad H_w = 10\ m \qquad r_0 = 0.35\ m$$

$$\text{Example 1} : \quad T_{w,I} = T_1 \ , \ T_{w,II} = T_3 \tag{45}$$

$$\text{Example 2} : \quad T_{w,I} = T_{w,II} = T_4$$

The optimal number of wells is

$$n_w = n_{w,I} + n_{w,II} = \frac{B_I}{\Delta w_I} + \frac{B_{II}}{\Delta w_{II}} \tag{46}$$

with $B_{I,II}$ - length of drainage contour (1500 m)

             $\Delta w_{I,II}$ - optimal well distance
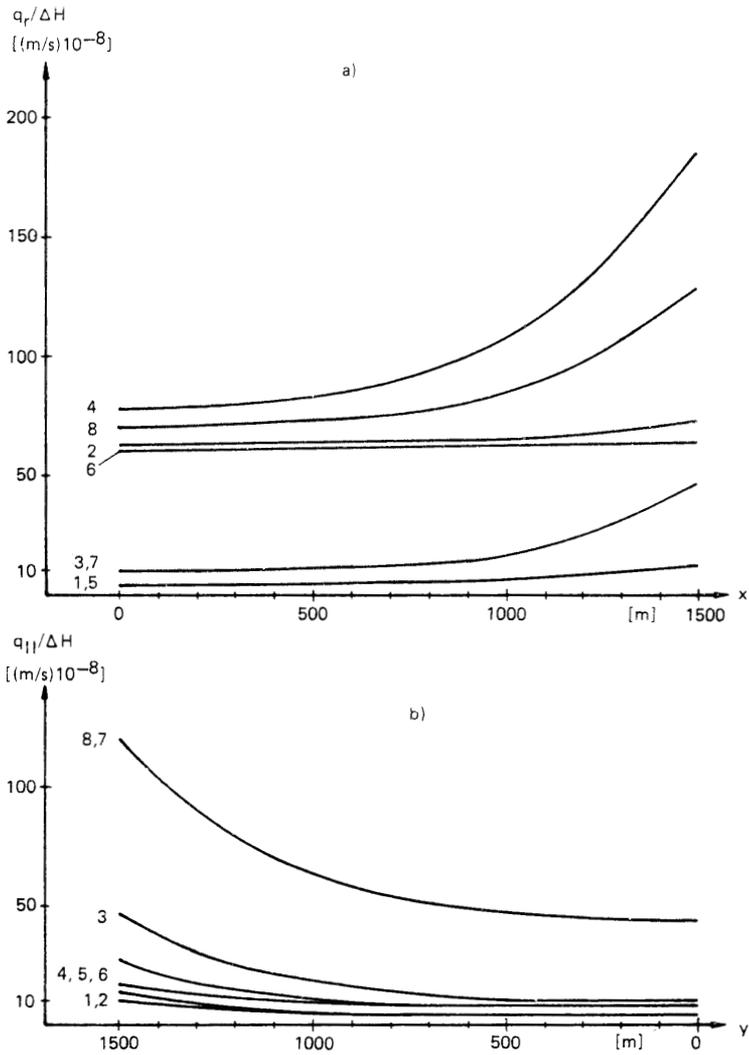
In Table 2 the results are summerized.

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Figure 7: Specific pumpage for hypothetical models
        a) contour I b) contour II

Based on Eq. (18) and (19) the effects of the parameters $T_{1,..,4}$ on the design param-
eter n_w have been estimated. In Table 3 selected results are listed.

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Table 2 : Drainage design for the hypothetical models

Number		1	2	3	4	5	6	7	8	9	10	11	12	13	14	15	16
Ex. 1	$n_{w,I}$	3	3	6	4	3	3	6	3								
	$n_{w,II}$	3	3	9	5	1	1	3	3								
	n_w	6	6	15	9	4	4	9	6								
Ex. 2	$n_{w,I}$	3	14	6	20	3	15	6	17	1	3	1	4	1	3	1	3
	$n_{w,II}$	3	2	5	4	4	4	15	15	1	1	1	1	1	1	3	3
	n_w	6	16	11	24	7	19	21	32	2	4	2	5	2	4	4	6

Table 3 : Effects of the parameters $T_{1,..,4}$ on the design parameter n_w

a) Main effects

		$\overline{n}_w$	Δn_w^1	Δn_w^2	Δn_w^3	Δn_w^4
Ex. 1	$n_{w,I}$	6.31	7.125	3.750	-0.375	-8.375
	$n_{w,II}$	4.00	−0.250	1.875	3.500	−5.000
	n_w	10.31	6.875	5.625	3.125	−13.375
Ex. 2	$n_{w,I}$	3.88	−1.25	1.75	−0.25	
	$n_{w,II}$	3.50	−1.00	2.50	−3.00	
	n_w	7.38	−2.25	4.25	−3.25	

$$\text{e.g.} \quad \Delta n_w^1 = \frac{1}{8}(16+24+19+32+4+5+4+6)-(6+11+7+21+2+2+2+4)=6.875$$

b) Interaction effects

		$\Delta n_w^{1,2}$	$\Delta n_w^{1,3}$	$\Delta n_w^{1,4}$	$\Delta n_w^{2,3}$	$\Delta n_w^{2,4}$	$\Delta n_w^{3,4}$
Ex. 1	$n_{w,I}$	-1.25	-0.25		-0.50		
	$n_{w,II}$	-1.00	-1.00		-1.50		
	n_w	-2.25	-1.25		-2.00		
Ex. 2	$n_{w,I}$	0.375	-0.375	-4.875	-0.625	-1.625	0.125
	$n_{w,II}$	0.000	0.250	0.250	2.750	-2.750	-2.500
	n_w	0.375	-0.125	-4.625	2.125	-4.375	-2.375

$$\text{e.g.} \quad \Delta n_w^{1,2} = \frac{1}{2}(\frac{1}{4}(24+32+5+6-11-21-2-4)-\frac{1}{4}(16+19+4+4-6-7-2-2))$$

$$\Delta n_w^{1,2} = \frac{1}{2}(7.25-6.50) = 0.375$$

For our example we assume $l = 3$ and $\alpha = 0.001$ as well as $\alpha = 0.01$. The specific cost are $c_w = 1.0$ and $c_b = 0.2$.

From Eq. (39) and (40) we obtain the minimum number of bore holes for T_4. The calculations are given in Table 4. The further calculations will be demonstrated for Example 2, subarea 4 (T_4) with $\alpha = 0.01$.

The optimal value of n_{b_4} is estimated based on Eq. (41) with $|\Delta n_w^4| = 13.375$. The optimal value of $n_{b,4}$ is estimated based on Eq. (41) with $|\Delta n_w^4| = 13.375$.

$$f(n_{b,4}) = \frac{13.375 \cdot 1.0}{2 \cdot 3} \cdot \frac{t_{a,m_4}}{\sqrt{n_w^4}} + 0.2 \cdot n_w^4 + 0.5 => Min. \tag{47}$$

with $\alpha = 0.01$. In Table 5 the results are depicted.

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Table 4: Calculation of the minimum number of bore holes

$n_{b,4}$	$t_{0.01,m_i}$	$\left[2 \cdot \dfrac{t_{0.01,m_i}}{3}\right]^2$	$t_{0.001,m_i}$	$\left[2 \cdot \dfrac{t_{0.001,m_i}}{3}\right]^2$
5	3.75	6.25		
6	3.37	5.05		
7	3.14	4.38		
8	3.0		4.79	11.06
9	2.90		4.50	9.0
10	2.82		4.30	8.22
11	2.76		4.14	7.62

$$\alpha = 0.01 \rightarrow \min n_i = 6 \quad ; \quad \alpha = 0.001 \rightarrow \min n_{b,i} = 9$$

Table 5: Calculation of the optimal number of bore holes for T_4

$n_{b,4}$	m_4	t_{a,m_4}	$f(n_{b,4})$
7	6	3.14	4.5
8	7	3.0	4.46
9	8	2.9	4.445
10	9	2.82	4.48
11	10	2.76	4.57

Table 6 shows the results for all subareas in the case of exploration in one step.

Table 6: Optimal number of bore holes for all subareas

Subarea (Parameter)		Example 1				Example 2			
		min V_i		opt n_b		min V_i		opt n_b	
		α =0.01	α =0.001	α =0.01	α =0.001	α =0.01	α =0.001	α =0.01	α =0.001
4	T_4	-	-	-	-	4.45	5.48	9	11
1	T_1	2.13	2.60	0	0	3.26	4.02	7	0
2	T_2	2.79	3.43	6	0	2.99	2.69	6	0
3	T_3	2.41	2.96	0	0	2.39	2.94	0	0

Finally some thoughts on the estimation of *exploration profitability* by the help of the calculated effects (Table 3). If we arrange the effects according to their values for n_w (Table 7), we get a general impression of the importance of different subareas for exploration.

For example 1 the exploration profitability decreases in the sequence T_2, T_3, T_1. The values are in the same range and the interaction effects are small. Consequently for Example 1 the subareas could be explored together.

In Example 2 the main effect decrease from T_4, T_1, T_2 up to T_3. Both, the high main effect for T_4, and the high interaction effects for T_1, T_2 indicates that subarea 4 should be explored first. This also becomes obvious looking at Table 6. After that should be decided on the exploration of the other subareas.

Table 7 : Selection of main effects and interaction effects

Main effects				Interaction effects	
Ex. 1	T_2	Δn_w^2	4.17	$\Delta n_w^{2,3}$	-2.00
	T_3	Δn_w^3	-3.25	$\Delta n_w^{1,2}$	-2.25
	T_1	Δn_w^1	-2.25	$\Delta n_w^{1,3}$	-1.25
Ex. 2	T_4	Δn_w^4	-13.375	$\Delta n_w^{1,4}$	-4.625
	T_1 ,	Δn_w^1	6.875	$\Delta n_w^{2,4}$	-4.375
	T_2	Δn_w^2	5.625	$\Delta n_w^{2,1}$	0.375
				$\Delta n_w^{3,4}$	-2.375
	T_3	Δn_w^3	3.125	$\Delta n_w^{3,1}$	-0.125
				$\Delta n_w^{3,2}$	2.125

In the case of step-wise exploration, the computed results for the hypothetical models (e.g. Figure 7) could be interpolated for the estimated parameter for the first explored subarea. The next working steps remain the same as before - with reduced number of hypothetical models.

CONCLUDING REMARKS

The proposed approach is an attempt to objectify the decision process in designing exploration programs. Based on calculations with hypothetical hydrogeological model the precision of planned exploration is economically evaluated.

The optimal exploration precision (number of research drillings) is estimated by minimizing the sum of cost of exploration and hypothetical losses, depending on the exploration precision. Although the method has been described with special regard to mine drainage systems the approach is applicable to other problems too, e.g. for the exploration for groundwater extraction (water work), for environmental protection measures needed due to mine drainage, etc. In general the hypothetical losses of all activities depending on exploration precision have to be summarized. The profitability of the exploration may be estimated, as a basis for step-wise exploration.

The demonstrated approach can be used for different parameter models, if the relationship between amount and precision of exploration can be quantified. The method is applicable for different goals of exploration, if the economic effect of the exploration can be quantified.

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